# Mathematics Involving Non-Parametric Statistics 

Yuk Tung Liu

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These short notes provide mathematical proofs of the following properties involving non-parametric statistics covered in Stat 200.

Wilcoxon Rank Sum: When there are no ties in both groups A and B, the expected value and variance of the rank sum for group A are

$$
\begin{equation*}
E\left(R_{A}\right)=\frac{n_{A}(N+1)}{2} \quad, \quad V\left(R_{A}\right)=\frac{n_{A} n_{B}(N+1)}{12}, \tag{0.1}
\end{equation*}
$$

where $n_{A}$ and $n_{B}$ are the number of observations in group A and B , respectively. The total number of observations $N=n_{A}+n_{B}$.

Relationship Between Wilcoxon Rank Sum and U Statistic: The Wilcoxon Rank Sum for group A, $R_{A}$, is related to the U statistic for group $\mathrm{A}, U_{A}$, by

$$
\begin{equation*}
R_{A}=U_{A}+\frac{n_{A}\left(n_{A}+1\right)}{2} \tag{0.2}
\end{equation*}
$$

Spearman's Rank-Order Correlation Coefficient: Suppose $x=\left(x_{1}, x_{2}, x_{3}, \cdots, x_{n}\right)$ and $y=\left(y_{1}, y_{2}, \cdots, y_{n}\right)$. If there are no ties in both $x$ and $y$, and $x$ and $y$ are uncorrelated, the expected value and variance of Spearman's rank-order correlation coefficient $r_{s}$ are

$$
\begin{equation*}
E\left(r_{s}\right)=0 \quad, \quad V\left(r_{s}\right)=\frac{1}{\sqrt{n-1}} \tag{0.3}
\end{equation*}
$$

## 1 Wilcoxon Rank Sum

To calculate the mean and variance of the rank sum $R_{A}$, we need to calculate several quantities.
We first want to calculate the values of two series: $1+2+3+\cdots+n$ and $1^{2}+2^{2}+\cdots+n^{2}$. The first one is an arithmetic series, which can be computed as follows.

$$
S_{n}=1+2+3+\cdots n
$$

We can rewrite $S_{n}$ as

$$
S_{n}=n+(n-1)+(n-2)+\cdots+1
$$

Adding the two expressions gives

$$
2 S_{n}=\underbrace{(n+1)+(n+1)+\cdots+(n+1)}_{n \text { times }}=n(n+1)
$$

Hence,

$$
\begin{equation*}
1+2+3+\cdots+n=\sum_{i=1}^{n} i=\frac{n(n+1)}{2} \tag{1.1}
\end{equation*}
$$

The second series is a discrete version of the integral

$$
\int_{1}^{n} x^{2} d x
$$

To calculate the sum, we consider the integral

$$
\int_{i-\frac{1}{2}}^{i+\frac{1}{2}} x^{2} d x=\frac{1}{3}\left[\left(i+\frac{1}{2}\right)^{3}-\left(i-\frac{1}{2}\right)^{3}\right]
$$

Now,

$$
\begin{align*}
\frac{1}{3}\left[\left(i+\frac{1}{2}\right)^{3}-\left(i-\frac{1}{2}\right)^{3}\right] & =\frac{1}{3}\left[\left(i^{3}+\frac{3}{2} i^{2}+\frac{3}{4} i+\frac{1}{8}\right)-\left(i^{3}-\frac{3}{2} i^{2}+\frac{3}{4} i-\frac{1}{8}\right)\right] \\
& =i^{2}+\frac{1}{12} \\
\Rightarrow & i^{2}=-\frac{1}{12}+\frac{1}{3}\left[\left(i+\frac{1}{2}\right)^{3}-\left(i-\frac{1}{2}\right)^{3}\right] \\
\Rightarrow \sum_{i=1}^{n} i^{2} & =-\frac{n}{12}+\frac{1}{3} \sum_{i=1}^{n}\left[\left(i+\frac{1}{2}\right)^{3}-\left(i-\frac{1}{2}\right)^{3}\right] \\
& =-\frac{n}{12}+\frac{1}{3} \sum_{i=1}^{n}\left(i+\frac{1}{2}\right)^{3}-\frac{1}{3} \sum_{i=1}^{n}\left(i-\frac{1}{2}\right)^{3} \\
& =-\frac{n}{12}+\frac{1}{3} \sum_{i=1}^{n}\left(i+\frac{1}{2}\right)^{3}-\frac{1}{3}\left(\frac{1}{2}\right)^{3}-\frac{1}{3} \sum_{i=2}^{n}\left(i-\frac{1}{2}\right)^{3} \\
& =-\frac{n}{12}+\frac{1}{3} \sum_{i=1}^{n}\left(i+\frac{1}{2}\right)^{3}-\frac{1}{24}-\frac{1}{3} \sum_{i=1}^{n-1}\left(i+\frac{1}{2}\right)^{3} \\
& =-\frac{n}{12}+\frac{1}{3}\left(n+\frac{1}{2}\right)^{3}-\frac{1}{24} \\
& =-\frac{n}{12}+\frac{1}{3}\left(n^{3}+\frac{3}{2} n^{2}+\frac{3}{4} n+\frac{1}{8}\right)-\frac{1}{24} \\
& =\frac{1}{3} n^{3}+\frac{1}{2} n^{2}+\frac{1}{6} n \\
& =\frac{n\left(2 n^{2}+3 n+1\right)}{6} \\
= & \frac{n(n+1)(2 n+1)}{6} \tag{1.2}
\end{align*}
$$

Let $x=$ permutation of $(1,2,3, \cdots, n)$. The expected value of $x_{i}$ is

$$
\begin{equation*}
E\left(x_{i}\right)=\frac{1}{n} \sum_{i=1}^{n} i=\frac{1}{n} \frac{n(n+1)}{2}=\frac{n+1}{2} \tag{1.3}
\end{equation*}
$$

The variance of $x_{i}$ is

$$
V\left(x_{i}\right)=E\left(x_{i}^{2}\right)-E^{2}\left(x_{i}\right)
$$

$$
\begin{align*}
& =\frac{1}{n} \sum_{i=1}^{n} i^{2}-\left(\frac{n+1}{2}\right)^{2} \\
& =\frac{(n+1)(2 n+1)}{6}-\frac{(n+1)^{2}}{4} \\
& =\frac{n^{2}-1}{12} \tag{1.4}
\end{align*}
$$

where we have used (1.2) in the third line.
The covariance $\operatorname{cov}\left(x_{i}, x_{j}\right)$ for $i \neq j$ can be computed as follows.

$$
\begin{align*}
\operatorname{cov}\left(x_{i}, x_{j}\right) & =E\left(x_{i} x_{j}\right)-E\left(x_{i}\right) E\left(x_{j}\right) \\
& =\frac{1}{n(n-1)} \sum_{i \neq j} i j-\left(\frac{n+1}{2}\right)^{2} \\
& =\frac{1}{n(n-1)} \sum_{i, j} i j-\frac{1}{n(n-1)} \sum_{i=1}^{n} i^{2}-\frac{(n+1)^{2}}{4} \\
& =\frac{1}{n(n-1)}\left(\sum_{i=1}^{n} i\right)^{2}-\frac{1}{n(n-1)} \frac{n(n+1)(2 n+1)}{6}-\frac{(n+1)^{2}}{4} \\
& =\frac{1}{n(n-1)} \frac{n^{2}(n+1)^{2}}{4}-\frac{(n+1)(2 n+1)}{6(n-1)}-\frac{(n+1)^{2}}{4} \\
& =-\frac{n+1}{12} \tag{1.5}
\end{align*}
$$

after straightforward algebra.
We are now ready to calculate the expected value and variance of Wilcoxon rank sum. Let $x_{1}, x_{2}, \cdots, x_{n_{A}}$ be the ranks of the elements of group A , and $x_{n_{A}+1}, x_{n_{A}+2}, \cdots, x_{N}$ be the ranks of the elements of group B . In the absence of ties, $x_{1}, x_{2}, \cdots, x_{N}$ is a permutation of $1,2,3, \cdots, N$. The rank sum for group A is

$$
\begin{gather*}
R_{A}=\sum_{i=1}^{n_{A}} x_{i} \\
\Rightarrow E\left(R_{A}\right)=\sum_{i=1}^{n_{A}} E\left(x_{i}\right)=\sum_{i=1}^{n_{A}} \frac{N+1}{2}=\frac{n_{A}(N+1)}{2} \tag{1.6}
\end{gather*}
$$

The variance can be computed as follows.

$$
\begin{aligned}
V\left(R_{A}\right) & =V\left(\sum_{i=1}^{n_{A}} x_{i}\right) \\
& =\sum_{i=1}^{n_{A}} V\left(x_{i}\right)+\sum_{i \neq j} \operatorname{cov}\left(x_{i}, x_{j}\right)
\end{aligned}
$$

Using (1.4) and (1.5), we have

$$
V\left(R_{A}\right)=\frac{n_{A}(N+1)}{12}-\frac{N+1}{12} \underbrace{\sum_{i \neq j} 1}_{=n_{A}\left(n_{A}-1\right)}
$$

$$
\begin{align*}
& =\frac{n_{A}(N+1)}{12}-\frac{n_{A}\left(n_{A}-1\right)(N+1)}{12} \\
& =\frac{n_{A}(N+1)\left(N-n_{A}\right)}{12} \\
& =\frac{n_{A} n_{B}(N+1)}{12} \tag{1.7}
\end{align*}
$$

## 2 Wilcoxon Rank Sum and U Statistic

Let $R_{A}$ be the Wilcoxon Rank Sum for group A and $U_{A}$ be the U statistic for group A. Then

$$
\begin{equation*}
R_{A}=U_{A}+\frac{n_{A}\left(n_{A}+1\right)}{2} \tag{2.1}
\end{equation*}
$$

The proof is very easy if there are no ties. It requires more algebra when there are ties. In the following, we first consider the case when there are no ties (not in group A at least). Then we tackle the general case when there are ties.

Case 1: No ties.
Sort the numbers in group A in ascending order: $A_{1}, A_{2}, \cdots, A_{n_{A}}$. Here $A_{i}\left(i=1,2, \cdots, n_{A}\right)$ denote the $i$ th number in group A when sorted in ascending order. Let $B_{1}, B_{2}, \cdots, B_{n_{B}}$ be the numbers in group B sorted in ascending order. By no ties we mean that $A_{i} \neq A_{j}$ for all $i \neq j\left[i, j \in\left(1, n_{A}\right)\right]$ and $A_{i} \neq B_{j}$ for all $i \in\left(1, n_{A}\right)$ and $j \in\left(1, n_{B}\right)$. In other words, all numbers in group A $A_{1}, A_{2}, \cdots, A_{n_{A}}$ are unequal and none of the numbers in group $\mathrm{B} B_{1}, B_{2}, \cdots, B_{n_{B}}$ is equal to any number in group A (However, there could be numbers in group B that are equal).

Let $r_{i}\left(i=1,2, \cdots, n_{A}\right)$ be the rank of $A_{i}$. Since there are no other numbers equal to $A_{i}, r_{i}$ is equal to one plus the number of numbers smaller than $A_{i}$. By definition, there are $i-1$ numbers in group A smaller than $A_{i}$ and $u_{i}$ numbers in group B smaller than $A_{i}$, where $u_{i}$ is the U count. Hence we have

$$
\begin{equation*}
r_{i}=i+u_{i} . \tag{2.2}
\end{equation*}
$$

Summing $i$ from 1 to $n_{A}$ gives

$$
\begin{equation*}
R_{A}=\sum_{i=1}^{n_{A}} r_{i}=\sum_{i=1}^{n_{A}} i+\sum_{i=1}^{n_{A}} u_{i}=\frac{n_{A}\left(n_{A}+1\right)}{2}+U_{A} \tag{2.3}
\end{equation*}
$$

Case 2: There are ties, meaning that at least two numbers in $A_{1}, A_{2}, \cdots, A_{n_{A}}, B_{1}, B_{2}, \cdots, B_{n_{B}}$ are equal.
There can also be multiple sets of ties. That is, at least two numbers are equal to a particular number, say $t_{1}$; at least two numbers are equal to another particular number, say $t_{2}$, ..etc. We can look at one particular set of ties, say $t_{q}$. Suppose there are $p$ numbers of $t_{q}$ occurring in groups A and B. Of these $p$ ties, $k$ of them are in group A and $p-k$ of them are in group B, where $k$ can be any integer between 0 and $p$. If $k=0$ (i.e. all the $p$ ties are in group B), we can disregard it because it has no effect on the rank sum $R_{A}$. So we focus on the case when $k$ is between 1 and $p$. Suppose the $k$ ties in group A are $A_{j}, A_{j+1}, \cdots A_{j+k-1}$. These $k$ ties have the same rank $r_{j}=r_{j+1}=\cdots=r_{j+k-1} \equiv r_{t_{q}}$ and the same U count $u_{j}=u_{j+1}=\cdots=u_{j+k-1} \equiv u_{t_{q}}$. From the definition of the rank in the case of ties, $r_{t_{q}}$ is equal to one plus the number of numbers smaller than $t_{q}$ plus $(p-1) / 2$. Now there are $j-1$ numbers smaller than $t_{q}$ in group A. Let $u_{j}^{\prime}$ be the number of numbers smaller than $t_{q}$ in group B. Then

$$
\begin{equation*}
r_{j}=r_{j+1}=\cdots=r_{j+k-1}=r_{t_{q}}=j+u_{j}^{\prime}+\frac{p-1}{2} . \tag{2.4}
\end{equation*}
$$

Recall that when there are ties, $1 / 2$ is contributed to the U count for the numbers in group B that are equal to $t_{q}$. Since there are $p-k$ numbers in group B equal to $t_{q}$, the U count is given by

$$
\begin{equation*}
u_{j}=u_{j+1}=\cdots=u_{j+k-1}=u_{t_{q}}=u_{j}^{\prime}+\frac{p-k}{2} . \tag{2.5}
\end{equation*}
$$

Combining equations (2.4) and (2.5) gives

$$
\begin{equation*}
r_{j}=r_{j+1}=\cdots=r_{j+k-1}=j+u_{j}-\frac{p-k}{2}+\frac{p-1}{2}=j+u_{j}+\frac{k-1}{2} . \tag{2.6}
\end{equation*}
$$

It follows that

$$
\begin{equation*}
\sum_{i=j}^{j+k-1} r_{i}=r_{j}+r_{j+1}+\cdots+r_{j+k-1}=k j+\frac{k(k-1)}{2}+k u_{j} . \tag{2.7}
\end{equation*}
$$

Since

$$
\begin{equation*}
\sum_{i=j}^{j+k-1} i=j+(j+1)+(j+2)+\cdots+(j+k-1)=k j+(1+2+\cdots k-1)=k j+\frac{k(k-1)}{2} \tag{2.8}
\end{equation*}
$$

we can write

$$
\begin{equation*}
\sum_{i=j}^{j+k-1} r_{i}=\sum_{i=j}^{j+k-1} i+\sum_{i=j}^{j+k-1} u_{i}=\sum_{i=j}^{j+k-1}\left(i+u_{i}\right) . \tag{2.9}
\end{equation*}
$$

This is the equation for a particular set of ties. The corresponding equations for the other sets of ties are equal to the equation above by changing $j$ and $k$ appropriate to the sets of ties. Summing over all ties appearing in group A, we have

$$
\begin{equation*}
\sum_{i \in \text { ties }} r_{i}=\sum_{i \in \text { ties }}\left(i+u_{i}\right) . \tag{2.10}
\end{equation*}
$$

When $A_{i}$ does not belong to any set of ties, equation (2.2) holds. Summing over numbers that don't belong to any set of ties, we have

$$
\begin{equation*}
\sum_{i \in \text { no ties }} r_{i}=\sum_{i \in \text { no ties }}\left(i+u_{i}\right) . \tag{2.11}
\end{equation*}
$$

Combining these two equations yield

$$
\begin{equation*}
\sum_{i=1}^{n_{A}} r_{i}=\sum_{i=1}^{n_{A}}\left(i+u_{i}\right)=\sum_{i=1}^{n_{A}} i+\sum_{i=1}^{n_{A}} u_{i}=\frac{n_{A}\left(n_{A}+1\right)}{2}+U_{A} . \tag{2.12}
\end{equation*}
$$

## 3 Spearman's Rank-Order Correlation Coefficient

Let $x=\left(x_{1}, x_{2}, \cdots, x_{n}\right)$ and $y=\left(y_{1}, y_{2}, \cdots, y_{n}\right)$. If there are no ties in both $x$ and $y$, we can replace $x$ by a permutation of $(1,2,3, \cdots, n)$ and replace $y$ by another permutation of $(1,2,3, \cdots, n)$. If $x$ and $y$ are uncorrelated, we have $E\left(f\left(x_{i}\right) g\left(y_{j}\right)\right)=E\left(f\left(x_{i}\right)\right) E\left(g\left(y_{j}\right)\right)$, where $f$ and $g$ are arbitrary functions.

The Spearman's rank-order correlation coefficient is

$$
\begin{equation*}
r_{s}=\frac{1}{n} \sum_{i=1}^{n} Z_{x_{i}} Z_{y_{i}}, \tag{3.1}
\end{equation*}
$$

where

$$
Z_{x_{i}}=\frac{x_{i}-\bar{x}}{S D_{x}} \quad, \quad Z_{y_{i}}=\frac{y_{i}-\bar{y}}{S D_{y}}
$$

are the Z-scores associated with $x$ and $y$. The mean and standard deviation of $x$ are

$$
\begin{equation*}
\bar{x}=\frac{1}{n} \sum_{i=1}^{n} x_{i} \quad, \quad S D_{x}=\sqrt{\frac{1}{n} \sum_{i=1}^{n}\left(x-x_{i}\right)^{2}}=\sqrt{\frac{1}{n} \sum_{i=1}^{n} x_{i}^{2}-\bar{x}^{2}} \tag{3.2}
\end{equation*}
$$

Since $x$ is a permutation of $(1,2,3, \cdots, n)$,

$$
\begin{equation*}
\bar{x}=\frac{1}{n} \sum_{i=1}^{n} i=\frac{n+1}{2} \quad, \quad S D_{x}=\sqrt{\frac{1}{n} \sum_{i=1}^{n} i^{2}-\frac{(n+1)^{2}}{4}}=\sqrt{\frac{n^{2}-1}{12}} \tag{3.3}
\end{equation*}
$$

using the results (1.1) and (1.2). Similarly,

$$
\begin{equation*}
\bar{y}=\bar{x}=\frac{n+1}{2} \quad, \quad S D_{y}=S D_{x}=\sqrt{\frac{n^{2}-1}{12}} . \tag{3.4}
\end{equation*}
$$

By construction, the expected value and variance of the Z-score are $E\left(Z_{x_{i}}\right)=E\left(Z_{y_{i}}\right)=0$ and $V\left(Z_{x_{i}}\right)=V\left(Z_{y_{i}}\right)=1$.
The expected value of $r_{s}$ is

$$
E\left(r_{s}\right)=E\left(\frac{1}{n} \sum_{i=1}^{n} Z_{x_{i}} Z_{y_{i}}\right)=\frac{1}{n} \sum_{i=1}^{n} E\left(Z_{x_{i}} Z_{y_{i}}\right)=\frac{1}{n} \sum_{i=1}^{n} E\left(Z_{x_{i}}\right) E\left(Z_{y_{i}}\right)=0
$$

The variance can be calculated as follows.

$$
\begin{align*}
V\left(r_{s}\right) & =\frac{1}{n^{2}} V\left(\sum_{i=1}^{n} Z_{x_{i}} Z_{y_{i}}\right) \\
= & \frac{1}{n^{2}}\left[\sum_{i=1}^{n} V\left(Z_{x_{i}} Z_{y_{i}}\right)+\sum_{i \neq j} \operatorname{cov}\left(Z_{x_{i}} Z_{y_{i}}, Z_{x_{j}} Z_{y_{j}}\right)\right]  \tag{3.5}\\
V\left(Z_{x_{i}} Z_{y_{i}}\right) & =E\left(Z_{x_{i}}^{2} Z_{y_{i}}^{2}\right)-E^{2}\left(Z_{x_{i}} Z_{y_{i}}\right) \\
& =E\left(Z_{x_{i}}^{2}\right) E\left(Z_{y_{i}}^{2}\right)-\left[E\left(Z_{x_{i}}\right) E\left(Z_{y_{i}}\right)\right]^{2} \\
& =E\left(Z_{x_{i}}^{2}\right) E\left(Z_{y_{i}}^{2}\right)
\end{align*}
$$

It follows from $V\left(Z_{x_{i}}\right)=V\left(Z_{y_{i}}\right)=1$ and $E\left(Z_{x_{i}}\right)=E\left(Z_{y_{i}}\right)=0$ that $V\left(Z_{x_{i}}\right)=E\left(Z_{x_{i}}^{2}\right)-E^{2}\left(Z_{x_{i}}\right)=E\left(Z_{x_{i}}^{2}\right)$. So $E\left(Z_{x_{i}}^{2}\right)=E\left(Z_{y_{i}}^{2}\right)=1$ and $V\left(Z_{x_{i}} Z_{y_{i}}\right)=1$. Thus,

$$
\begin{align*}
V\left(r_{s}\right)=\frac{1}{n^{2}} \sum_{i=1}^{n} 1+\frac{1}{n^{2}} & \sum_{i \neq j} \operatorname{cov}\left(Z_{x_{i}} Z_{y_{i}}, Z_{x_{j}} Z_{y_{j}}\right)=\frac{1}{n}+\frac{1}{n^{2}} \sum_{i \neq j} \operatorname{cov}\left(Z_{x_{i}} Z_{y_{i}}, Z_{x_{j}} Z_{y_{j}}\right)  \tag{3.6}\\
\operatorname{cov}\left(Z_{x_{i}} Z_{y_{i}}, Z_{x_{j}} Z_{y_{j}}\right) & =E\left(Z_{x_{i}} Z_{x_{j}} Z_{y_{i}} Z_{y_{j}}\right)-E\left(Z_{x_{i}} Z_{y_{i}}\right) E\left(Z_{x_{j}} Z_{y_{j}}\right) \\
& =E\left(Z_{x_{i}} Z_{x_{j}}\right) E\left(Z_{y_{i}} Z_{y_{j}}\right)-E\left(Z_{x_{i}}\right) E\left(Z_{y_{i}}\right) E\left(Z_{x_{j}}\right) E\left(Z_{y_{j}}\right) \\
& =E\left(Z_{x_{i}} Z_{x_{j}}\right) E\left(Z_{y_{i}} Z_{y_{j}}\right) \\
& =\operatorname{cov}\left(Z_{x_{i}}, Z_{x_{j}}\right) \operatorname{cov}\left(Z_{y_{i}}, Z_{y_{j}}\right) \\
& =\left[\operatorname{cov}\left(Z_{x_{i}}, Z_{x_{j}}\right)\right]^{2} \tag{3.7}
\end{align*}
$$

since $\operatorname{cov}\left(Z_{x_{i}}, Z_{x_{j}}\right)=\operatorname{cov}\left(Z_{y_{i}}, Z_{y_{j}}\right)$.

$$
\operatorname{cov}\left(Z_{x_{i}}, Z_{x_{j}}\right)=\operatorname{cov}\left(\frac{x_{i}-\bar{x}}{S D_{x}}, \frac{x_{j}-\bar{x}}{S D_{x}}\right)
$$

$$
\begin{align*}
& =\frac{1}{S D_{x}^{2}} \operatorname{cov}\left(x_{i}, x_{j}\right) \\
& =-\frac{12}{n^{2}-1} \frac{n+1}{12} \\
& =-\frac{1}{n-1} \tag{3.8}
\end{align*}
$$

Hence,

$$
\begin{align*}
V\left(r_{s}\right) & =\frac{1}{n}+\frac{1}{n^{2}} \frac{1}{(n-1)^{2}} \sum_{i \neq j} 1 \\
& =\frac{1}{n}+\frac{1}{n^{2}} \frac{1}{(n-1)^{2}} n(n-1) \\
& =\frac{1}{n}+\frac{1}{n(n-1)} \\
& =\frac{1}{n}\left(1+\frac{1}{n-1}\right) \\
& =\frac{1}{n} \frac{n}{n-1} \\
& =\frac{1}{n-1} \tag{3.9}
\end{align*}
$$

